

# Factors or Fake?

## A New Look at Anomalies

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*Any resemblance to real anomalies is purely coincidental . . . until we compute posterior probabilities.*

## Introduction

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- The asset-pricing literature has produced hundreds of candidate return anomalies, indexed by  $a_1, \dots, a_n$ .
- A growing body of work argues that *robust anomalies may be far fewer than previously believed*.
- Prominent contributions include Harvey, Liu, and Zhu (RFS 2016), Hou, Xue, and Zhang (RFS 2020), and Chordia, Goyal, and Saretto (RFS 2020).
- Our main is to revisit this screening question from a new perspective.

- Existing tests all follow the same recipe.
- Specifically, for each anomaly  $a_i$ , one sets up the hypotheses

$H_{0,i}$  :  $a_i$  is spanned by a benchmark SDF,

$H_{1,i}$  :  $a_i$  is not spanned by the benchmark SDF,

where the benchmark is, for example, CAPM, FF3, or FF6.

- Rejection of  $H_{0,i}$  is then interpreted as “discovery” (that  $a_i$  carries pricing information not captured by the benchmark SDF).

## Existing Testing Paradigm (contd)

- It has been recognized, however, that one should only make claims for discovery after accounting for
  - multiple testing and false discoveries,
  - data snooping and publication bias,
  - balancing of Type I and Type II errors.
- For example, [McLean and Pontiff \(JF 2016\)](#), [Linnainmaa and Roberts \(RFS 2018\)](#), [Giglio, Liao, and Xiu \(RFS 2021\)](#), ([Harvey and Liu, JF 2020](#); [Chen, MS 2025](#)).

## Is There a Problem?

- Despite many statistical advances, there is a *conceptual* issue arising from how the hypotheses are formulated.
- In the current paradigm, “discovery” means that  $a_i$  is not spanned by the chosen benchmark.
- Such a claim can be justified only if the chosen benchmark is the true SDF.
- Why? Because an anomaly that is not spanned by one benchmark may be spanned by another benchmark that is closer to the true SDF.

## Methodology

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## Reversing the Question

- We ask a different question: “can  $a_i$  be spanned by anything at all?” If  $a_i$  can be spanned by *any* admissible SDF, it cannot be an anomaly.
- This insight leads to a new pair of hypotheses:

$H_{0,i}$  :  $a_i$  is not spanned by any admissible SDF,

$H_{1,i}$  :  $a_i$  is spanned by at least one admissible SDF.

- This is a much tougher standard for an anomaly to survive.
- Importantly, evidence of spanning (the discovery) can be established without knowing the true SDF.

## Frequentist Test? Not easily possible

- In the usual case, suppose we fix a candidate SDF model  $M$ .
- Spanning reduces to a standard test of pricing error:

$$H_0 : \alpha = 0.$$

- The standard test statistic is  $t$  or  $F$ .
- Under  $H_0$ : central  $t$  or  $F$  distribution.
- Under the alternative: non-central distribution.
- For a *single* model, frequentist inference is well defined.

## Our Null is Not a Single-Model Null

- Our null is

$H_0$  :  $a_i$  is not spanned by any admissible model.

- This means every admissible model has nonzero pricing error.
- A frequentist test would require:
  - testing spanning under every admissible model,
  - correcting for the model search.
- The global statistic becomes the minimum of many tests:

$$T^* = \min_j T_j,$$

where each  $T_j$  is non-central under  $H_0$ .

## Why the Frequentist Route Breaks

- The null distribution of  $T^*$  depends on:
  - unknown noncentrality parameters,
  - dependence across models,
  - the size and geometry of the model space,
  - the search procedure itself.
- The test is therefore **non-pivotal** and not identifiable.
- There is no model-free sampling distribution.

- Bayes is natural here: it treats the hypotheses symmetrically and allows integration over the entire model space.
- We compute posterior probabilities of spanning by scanning all admissible SDF decompositions (Chib and Zeng, JBES 2020; Chib, Zeng and Zhao, JF 2020; Chib, Zhao and Zhou, MS 2024).
- We then apply Bayesian multiple testing with false discovery control (Müller, Parmigiani and Rice, 2007; Newton et al., 2004) to detect the *fake anomalies* among  $\{a_i\}$ .

## Anomaly Spanning: One at a Time

- $n$  candidate anomalies:  $a_1, a_2, \dots, a_n$
- For each anomaly  $a_i$ , we could simply compare the two models
  - FF6 unspanned,  $a_i$  spanned.

$$\text{FF6} = \boldsymbol{\lambda}^{(i)} + \boldsymbol{u}^{(i)},$$

$$a_i = \Gamma^{(i)} \boldsymbol{x}^{(i)} + \boldsymbol{\varepsilon}^{(i)}$$

- $a_i$  not in the span of FF6

$$\begin{pmatrix} \text{FF6} \\ a_i \end{pmatrix} = \boldsymbol{\lambda}^{(i)} + \boldsymbol{u}^{(i)}$$

- This comparison is incomplete, however. There are many distinct ways in which  $a_i$  can be spanned or not spanned by subsets of FF6.

## Enumerating Spanning Configurations

- Define the augmented factor set

$$\mathbf{f}^{(i)} = (\text{FF6}, a_i)$$

- Consider all possible splits of  $\mathbf{f}^{(i)}$  ( $j = 1, 2, \dots, J$ , for  $J = 127 = 2^7 - 1$ ):
  - factors that are unspanned  $\mathbf{x}_j^{(i)}$
  - factors that are spanned  $\mathbf{w}_j^{(i)}$
  - Each of these splits defines a model:  $\mathbb{M}_j^{(i)} = (\mathbf{x}_j^{(i)}, \mathbf{w}_j^{(i)})$

**Table 1: Candidate Splits for Factor Decomposition (Benchmark: FF6; one at a time).**

Model	$x$	$w$
$M_1^{(i)}$	Mkt	SMB, HML, RMW, CMA, MOM, $a_i$
$M_2^{(i)}$	SMB	Mkt, HML, RMW, CMA, MOM, $a_i$
$M_3^{(i)}$	HML	Mkt, SMB, RMW, CMA, MOM, $a_i$
$M_4^{(i)}$	RMW	Mkt, SMB, HML, CMA, MOM, $a_i$
$M_5^{(i)}$	CMA	Mkt, SMB, HML, RMW, MOM, $a_i$
$M_6^{(i)}$	MOM	Mkt, SMB, HML, RMW, CMA, $a_i$
$M_7^{(i)}$	$a_i$	Mkt, SMB, HML, RMW, CMA, MOM
$M_8^{(i)}$	Mkt, SMB	HML, RMW, CMA, MOM, $a_i$
$M_9^{(i)}$	Mkt, HML	SMB, RMW, CMA, MOM, $a_i$
...		
$M_{13}^{(i)}$	Mkt, $a_i$	SMB, HML, RMW, CMA, MOM
...		
$M_{120}^{(i)}$	Mkt, SMB, HML, RMW, CMA, MOM	$a_i$
...		
$M_{127}^{(i)}$	Mkt, SMB, HML, RMW, CMA, MOM, $a_i$	$\emptyset$

## Model-Level Spanning Decompositions

- For each anomaly  $a_i$ , each admissible decomposition  $\mathbb{M}_j^{(i)} = (\mathbf{x}_j^{(i)}, \mathbf{w}_j^{(i)})$  corresponds to the system

$$\mathbf{x}_j^{(i)} = \boldsymbol{\lambda}_j^{(i)} + \mathbf{u}_j^{(i)}, \quad (1)$$

$$\mathbf{w}_j^{(i)} = \Gamma_j^{(i)} \mathbf{x}_j^{(i)} + \boldsymbol{\varepsilon}_j^{(i)}, \quad j = 1, \dots, 127. \quad (2)$$

- Each model encodes a distinct hypothesis about which factors (including  $a_i$ ) are spanned and which are not.
- Using the Bayesian model scanning methods of [Chib and Zeng \(2020\)](#) and [Chib, Zeng, and Zhao \(2020\)](#), we compute the marginal likelihood and posterior probability of each model:

$$\Pr(\mathbb{M}_j^{(i)} \mid \text{data}) = \frac{m_j^{(i)}}{\sum_{\ell=1}^J m_{\ell}^{(i)}}, \quad j = 1, \dots, J, \quad i = 1, \dots, n.$$

## From Model Probabilities to Spanning Probabilities

- The posterior probabilities  $\Pr(\mathbb{M}_j^{(i)} \mid \text{data})$  summarize evidence across *all* admissible spanning decompositions.
- For each anomaly  $a_i$ , we aggregate this information to obtain the posterior probability that  $a_i$  is spanned by *some* subset of FF6:

$$p_i = \Pr(a_i \text{ is spanned} \mid \text{data}) = \sum_{j: a_i \in \mathbf{w}_j^{(i)}} \Pr(\mathbb{M}_j^{(i)} \mid \text{data}), \quad i = 1, \dots, n.$$

- Thus,  $p_i$  measures the total posterior mass assigned to models in which  $a_i$  is spanned.

## Bayesian EFDR-Based Selection

- The posterior probability  $p_i$  directly induces the *local false discovery rate* (lfdr):

$$\text{lfdr}_i = \Pr(H_{0,i} \mid \text{data}) = 1 - p_i,$$

where  $H_{0,i}$  denotes the hypothesis that  $a_i$  is *not* spanned.

- For any selected set  $S \subset \{1, \dots, n\}$  (interpreted as “spanned”), a false discovery occurs when  $i \in S$  but  $H_{0,i}$  is true.
- The posterior *expected false discovery rate* (EFDR) is

$$\text{EFDR}(S \mid \text{data}) = \frac{\sum_{i \in S} \text{lfdr}_i}{\max\{|S|, 1\}}.$$

- This criterion delivers a principled, multiplicity-aware selection of spanned anomalies.

● *EFDR: One-at-a-Time Selection:*

- Sort  $\text{lfdr}_{(1)} \leq \dots \leq \text{lfdr}_{(n)}$ .
- For  $k = 1, \dots, n$ , define

$$\text{EFDR}(k) = \frac{1}{k} \sum_{\ell=1}^k \text{lfdr}_{(\ell)} = \frac{1}{k} \sum_{\ell=1}^k (1 - p_{(\ell)}).$$

- Choose

$$k^* = \max\{k : \text{EFDR}(k) \leq q\},$$

and set

$$S_1^* = \{a_{(1)}, \dots, a_{(k^*)}\}.$$

- Exit the one-at-a-time testing with the unspanned set

$$A_2^* = \{a_1, \dots, a_n\} \setminus S_1^*$$

## Proposition (One-at-a-Time Anomaly Tests)

For individual anomaly tests  $\{a_i\}_{i=1}^n$ , a step-up selection rule based on posterior local false discovery rates (lfdr) satisfies the following:

- 1 It controls the expected false discovery rate (EFDR) of the selected set at level  $q$ .
- 2 Among all selection rules that satisfy this EFDR constraint, the step-up rule is *Bayes-optimal* in the sense of maximizing the expected number of true discoveries.

## Anomaly Spanning: Two at-a-Time Tests

- In practice, anomalies may be *jointly* spanned by the same SDF.
- We therefore extend the analysis to joint tests involving groups of anomalies.
- Let  $\mathcal{A}_2^* \subset \{1, \dots, n\}$  denote the set of anomalies remaining after the one-at-a-time screening stage.

## Candidate Groups of Fixed Size $s$

- In the second step, fix a group size of  $s = 2$ .
- Consider all distinct groups of size  $s$  drawn from  $\mathcal{A}_2^*$ .
- Let

$$\mathcal{G}_1, \dots, \mathcal{G}_{N_s}$$

denote these groups, where

$$N_s = \binom{|\mathcal{A}_2^*|}{s}.$$

- Each  $\mathcal{G}_i$  represents a candidate set of anomalies that may be *jointly* spanned.

## Joint Spanning Hypotheses

- For each group  $\mathcal{G}_i$ , we test the joint spanning hypotheses:

$H_{0,i}^{(s)}$  : none of the anomalies in  $\mathcal{G}_i$  are spanned,

$H_{1,i}^{(s)}$  : all anomalies in  $\mathcal{G}_i$  are jointly spanned.

- Let

$$p_{i,s}^{\text{joint}} = \Pr\left(H_{1,i}^{(s)} \mid \text{data}\right)$$

denote the posterior probability of joint spanning for group  $\mathcal{G}_i$ .

- The corresponding local false discovery rate is

$$\text{lfdr}_{i,s} = 1 - p_{i,s}^{\text{joint}}.$$

## Multiplicity and EFDR Control (Group Level)

- Pool the  $N_s$  groups  $\{\mathcal{G}_i\}_{i=1}^{N_s}$  and sort their lfdrs:

$$\text{lfdr}_{(1),s} \leq \text{lfdr}_{(2),s} \leq \cdots \leq \text{lfdr}_{(N_s),s}.$$

- For  $k = 1, \dots, N_s$ , define the expected false discovery rate:

$$\text{EFDR}_s(k) = \frac{1}{k} \sum_{\ell=1}^k \text{lfdr}_{(\ell),s}.$$

- Selection rule:** choose

$$k_s^* = \max\{k : \text{EFDR}_s(k) \leq q\}, \quad q \in (0, 1).$$

- Given  $k_s^*$ , the set of anomalies classified as spanned at stage  $s$  is

$$\mathcal{S}_s^* = \bigcup_{\ell=1}^{k_s^*} \mathcal{G}_{(\ell)}.$$

- These anomalies are removed from further testing before proceeding to larger group sizes.

### Proposition (Global EFDR Control)

Let  $\mathcal{S}_1^*, \dots, \mathcal{S}_S^*$  denote the subsets declared spanned at each stage, with increasing group sizes. Define the cumulative discovery set as

$$\mathcal{S}^* = \bigcup_{s=1}^S \mathcal{S}_s^*.$$

The posterior expected false discovery rate for the cumulative discovery set is

$$\text{EFDR}(\mathcal{S}^* \mid \mathcal{D}) = \frac{1}{\max(|\mathcal{S}^*|, 1)} \sum_{S \in \mathcal{S}^*} \text{lfdr}_S.$$

Under nested elimination, this global EFDR satisfies

$$\text{EFDR}(\mathcal{S}^* \mid \mathcal{D}) \leq q.$$

## Empirical Analysis

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*"When you have eliminated the impossible, whatever remains, however improbable, must be the truth." — Sherlock Holmes*

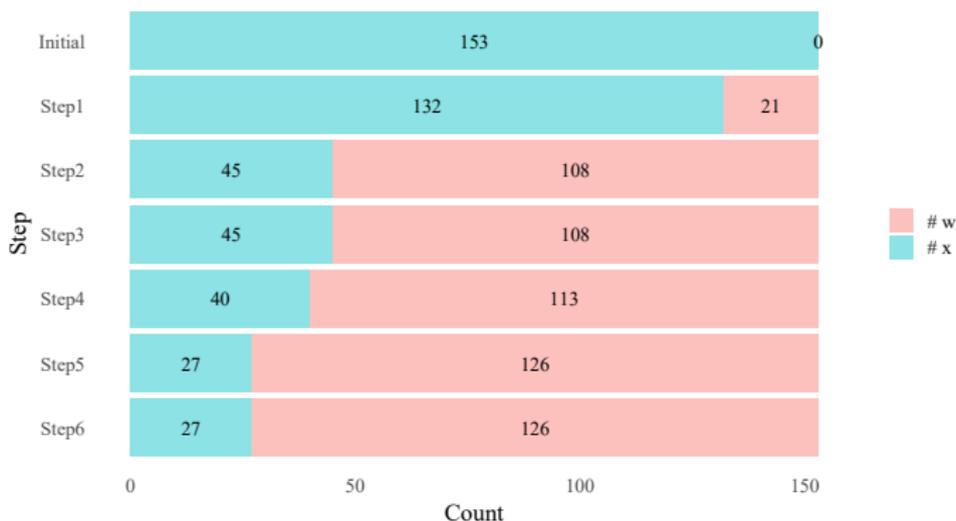
- We take the methodology to the data.
- The evidence shows that many anomalies are fake.
- The conclusions are robust across two benchmarks.
- An SDF built from the benchmark and the survivors improves both pricing fit and portfolio performance.

- 153 anomalies grouped into 13 thematic categories (Jensen, Kelly, and Pedersen, JF 2023)
- January 1985 through December 2024,  $T = 480$  months.
- Training sample: the first 25% of observations ( $T_{tr} = 120$ )
- **Test assets:**
  - Fama-French 49 industry portfolios
  - 100 P-Tree portfolios (Cong, Feng, He and He, JFE 2025)
  - 120 ( $2 \times 3 \times 20$ ) bivariate-sorted portfolios

MKT	Market risk: Compensation for aggregate market exposure; CAPM backbone.
SMB	Size: Small minus big firms; captures size-related frictions and financing constraints.
HML	Value: High vs. low book-to-market firms; classic value (cheap vs. expensive).
RMW	Profitability: Robust minus weak operating performance; often interpreted as a quality factor.
CMA	Investment: Conservative vs. aggressive investment; grounded in q-theory of investment.
UMD	Momentum: Return continuation; winners minus losers.

## Fake Anomalies: FF6 Benchmark

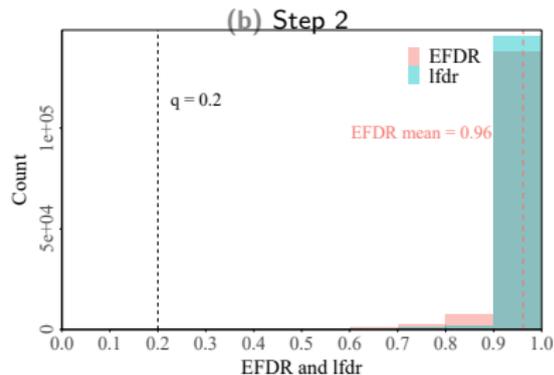
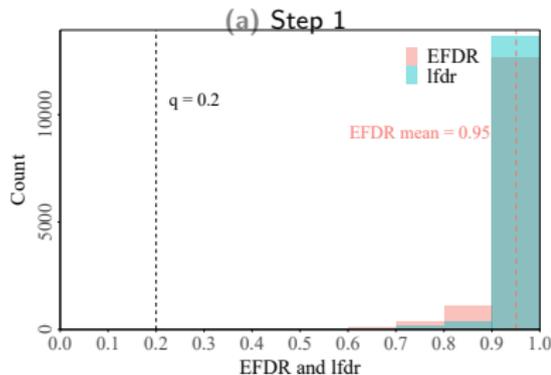
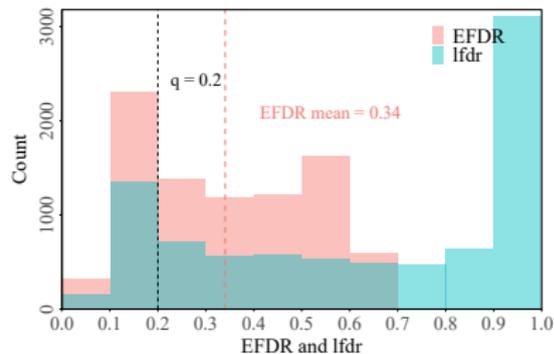
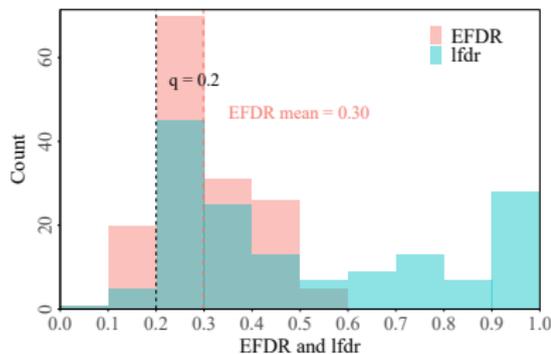
Figure 1: Counts of Fake Anomalies and Remaining Anomalies.



- 27 *last-stage anomalies* and 126 *fake anomalies*.
- Steps 1, 2, 4, and 5 contribute 21, 87, 5, and 13 new fake anomalies, respectively, while Step 3 yields none.

## (ii) Fake Anomalies: FF6 Benchmark

Figure 2: Histogram of EFDRs across Steps 1–4.



(c) Step 3

(d) Step 4

## Why These?

Table 2: Number and Percentage of Anomalies.

Theme	# <i>w</i>	# <i>x</i>	(%) <i>w</i>	(%) <i>x</i>
Investment	22	0	100.0%	0.0%
Low risk	18	0	100.0%	0.0%
Momentum	8	0	100.0%	0.0%
Profit growth	11	1	91.7%	8.3%
Debt issuance	6	1	85.7%	14.3%
Short-term reversal	5	1	83.3%	16.7%
Size	4	1	80.0%	20.0%
Seasonality	10	2	83.3%	16.7%
Value	15	3	83.3%	16.7%
Accruals	3	3	50.0%	50.0%
Low leverage	7	4	63.6%	36.4%
Profitability	7	4	63.6%	36.4%
Quality	10	7	58.8%	41.2%
Total	126	27	82.4%	17.6%

- The closer a factor's economic content is to the benchmark, the more likely it is to be spanned and thus classified as *w*.

## (ii) Evaluating Fake Anomalies and Last-stage Anomalies

- The last-stage anomalies are either genuine factors or not: we cannot ascertain since **the true SDF is unknown**.
- We can assert: up to the level of our EFDR control, 126 anomalies can be spanned by FF6 or by any proper subset of FF6  
⇒ **these can be classified as fake anomalies**.

## (ii) Evaluating Fake Anomalies and Last-stage Anomalies

- OOS predictive performance: predictive likelihood (Chib, Zeng and Zhao, MS 2024)
  - Sampling 1,000 instances of  $5x$  from  $27x$ , and 1,000 instances of  $5w$  from remaining  $w \Rightarrow 1,000,000$  pairs of  $(5x, 5w)$ .
  - Panel A: in each pair, consider two models ((i) “True”: risk factors are  $5x$ ; (ii) “Reverse”: risk factors are  $5w$ ) and calculate the predictive likelihoods.
  - Panel B: Risk factors are  $(5x, FF6)$  (“True”), or  $(5w, FF6)$  (“Reverse”).

**Table 3:** Statistics of Log Average of Predictive LLH for Factor Combinations

Factor Set		Mean	Std.	Median	2.5% qtile	97.5 % qtile
<i>Panel A</i>						
$(5x, 5w)$	True	3241.5	113.5	3236.5	3033.8	3474.5
	Reverse	3236.7	113.8	3231.8	3028.1	3470.1
<i>Panel B</i>						
$(FF6, 5x, 5w)$	True	5147.2	112.2	5143.3	4939.4	5375.0
	Reverse	5143.4	112.9	5139.7	4933.8	5372.3

### (iii) Evaluating Fake Anomalies and Last-stage Anomalies

Table 4: Average Sharpe Ratios of Tangency Portfolios Across Different Factors.

$f$	# Factor	# Comb	IS	OOS
<i>Panel A: Only Anom.</i>				
5x	5	80,730	0.99	0.86
5w	5	50,000	0.90	0.24
<i>Panel B: Anom. &amp; Bench. (Match with Method)</i>				
(5x, FF6)	5+6	80,730	2.14	1.18
(5w, FF6)	5+6	50,000	1.93	0.69
<i>Panel C: Anom. and Bench.</i>				
(27x, FF6)	27+6	1	3.09	1.01
(27w, FF6)	27+6	50,000	3.06	0.48

## (iv) Evaluating Fake Anomalies and Last-stage Anomalies

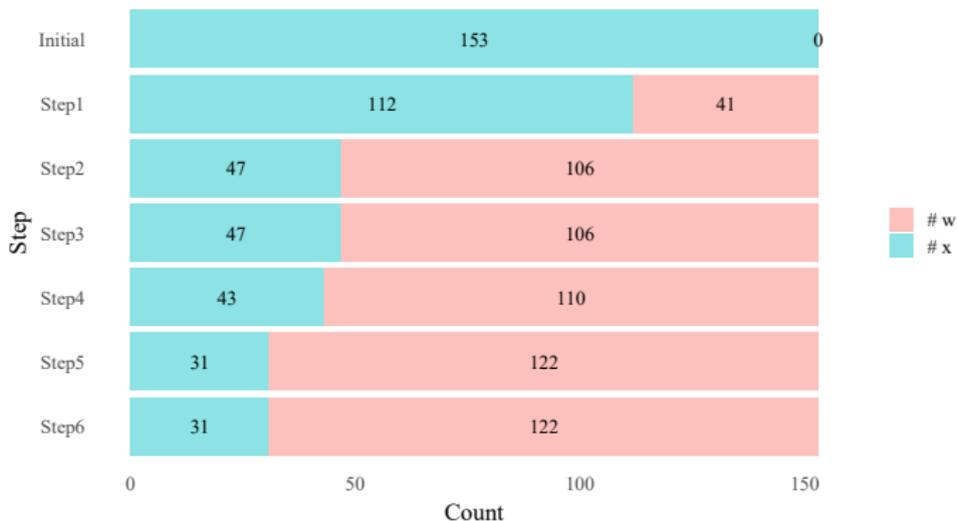
Table 5: Average Explanatory Power of Different Factors Across Test Assets.

$f$	# Factor	# Comb	Metrics				
			RMS $\alpha$	RMSE	$ \alpha $	GRS stat.	Total $R^2$
<i>Panel A. 49 Industry</i>							
(5x, FF6)	11	80,730	0.0041	0.0474	0.0032	2.10	53.2
(5w, FF6)	11	50,000	0.0043	0.0480	0.0033	2.33	52.1
<i>Panel B. 100 P-Tree</i>							
(5x, FF6)	11	80,730	0.0038	0.0387	0.0028	2.93	65.8
(5w, FF6)	11	50,000	0.0038	0.0387	0.0028	3.07	65.8
<i>Panel C. 120 Bisort</i>							
(5x, FF6)	11	80,730	0.0019	0.0149	0.0015	13.46	95.0
(5w, FF6)	11	50,000	0.0019	0.0152	0.0015	14.67	94.8

- NewB6: Retain Mkt but replace the other FF6 factors with:
  - the profitability (ROE) and investment (IA) factors from [Hou, Xue, and Zhang \(RFS, 2020\)](#) q-factor model
  - [Pástor and Stambaugh \(JPE, 2003\)](#) liquidity factor (LIQ)
  - [Frazzini and Pedersen \(JFE, 2014\)](#) betting-against-beta factor (BAB)
  - [Asness and Frazzini \(JPM, 2013\)](#) value factor (HMLD)

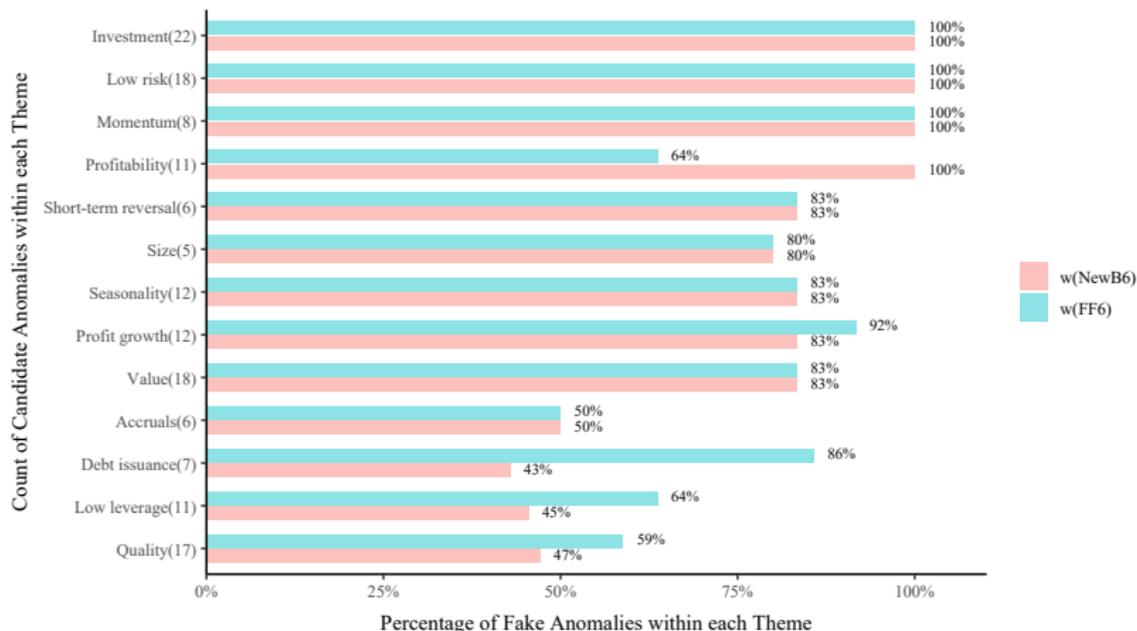
## (ii) Fake Anomalies: Another Benchmark

Figure 4: Counts of Fake Anomalies and Remaining Anomalies (Benchmark: NewB6).



# Fake Anomalies

Figure 5: Percentage of Fake Anomalies within each Theme.



- # fake anomalies: 126(FF6) and 122(NewB6), respectively.  
The intersection of  $w_{FF6}$  and  $w_{NewB6}$  contains 116 elements.  
⇒ Substantial overlap despite the two benchmarks sharing only Mkt.

## What Remains?

Table 6: 21 Remaining Anomalies Across Benchmarks.

Names of x	Theme	Description
cowc_gr1a	Accruals	Change in current operating working capital
oaccruals_at	Accruals	Operating accruals
oaccruals_ni	Accruals	Percent operating accruals
noa_at	Debt issuance	Net operating assets
cash_at	Low leverage	Cash-to-assets
netdebt_me	Low leverage	Net debt-to-price
rd_sale	Low leverage	R&D-to-sales
rd5_at	Low leverage	R&D capital-to-book assets
sale_emp_gr1	Profit growth	Labor force efficiency
cop_at	Quality	Cash-based operating profits-to-book assets
cop_at11	Quality	Cash-based operating profits-to-lagged book assets
gp_at	Quality	Gross profits-to-assets
qmj	Quality	Quality minus Junk: Composite
qmj_safety	Quality	Quality minus Junk: Safety
sale_bev	Quality	Assets turnover
seas_6_10an	Seasonality	Years 6-10 lagged returns, annual
seas_11_15an	Seasonality	Years 11-15 lagged returns, annual
rd_me	Size	R&D-to-market
rmax5_rvol_21d	Short-term reversal	Highest 5 days of return scaled by volatility
fcf_me	Value	Free cash flow-to-price
div12m_me	Value	Dividend yield

## (ii) What Remains?

- **Fundamental Quality:**

cash-based profitability (`cop_at`, `cop_at11`), gross profitability (`gp_at`), composite quality (`qmj`, `qmj_safety`), asset turnover (`sale_bev`), labor productivity (`sale_emp_gr1`), accrual quality (`cowc_gr1a`, `oaccruals_at`, `oaccruals_ni`), and balance-sheet defensiveness (`cash_at`, `netdebt_me`)

- **Management Decisions:**

net operating assets (`noa_at`); a suite of R&D intensity metrics (`rd_sale`, `rd5_at`, `rd_me`)

- **Valuation Level:**

cash-flow yields: free-cash-flow yield (`fcf_me`) and dividend yield (`div12m_me`).

- **Market Behavior and Patterns:**

short-horizon reversal (`rmax5_rvol_21d`), and long-lag seasonality (`seas_6_10an`, `seas_11_15an`).

$$\mathbf{x}^* = \underbrace{\boldsymbol{\lambda}_{x^*}}_{(6+21) \times 1} + \boldsymbol{\epsilon}_{x^*,t}, \quad \boldsymbol{\epsilon}_{x^*,t} \sim \mathcal{N}_{27}(\mathbf{0}, \boldsymbol{\Omega}_{x^*})$$

- Market prices of factor risks:  $\boldsymbol{\Omega}_{x^*}^{-1} \boldsymbol{\lambda}_{x^*}$

## Posterior summary of Market Prices of Factor Risks

Name	Theme	Post. mean	Post. std.	Post. median	2.5% qtile	97.5% qtile
Mkt		6.5	2.1	6.5	<b>2.4</b>	<b>10.7</b>
SMB		4.9	3.1	4.8	-1.3	11.1
HML		-0.1	4.1	-0.1	-8.2	8.0
RMW		18.4	5.5	18.3	<b>7.8</b>	<b>29.3</b>
CMA		16.5	5.1	16.5	<b>6.8</b>	<b>26.6</b>
MOM		4.8	2.0	4.8	<b>0.9</b>	<b>8.7</b>
cowc_gr1a	Accruals	1.5	5.0	1.5	-8.3	11.4
oaccruals_at	Accruals	-6.5	7.0	-6.5	-20.2	7.1
oaccruals_ni	Accruals	8.7	6.7	8.6	-4.3	21.8
noa_at	Debt issuance	3.6	5.2	3.6	-6.5	13.8
cash_at	Low leverage	5.8	6.3	5.8	-6.6	18.3
netdebt_me	Low leverage	5.9	6.2	5.9	-6.2	17.9
rd_sale	Low leverage	5.9	5.3	5.9	-4.5	16.3
rd5_at	Low leverage	-10.8	5.5	-10.7	-21.5	-0.2
sale_emp_gr1	Profit growth	-5.4	3.9	-5.4	-13.2	2.2
cop_at	Quality	-0.3	14.2	-0.3	-28.2	27.4
cop_at11	Quality	18.0	14.4	18.0	-10.1	46.3
gp_at	Quality	-23.7	8.2	-23.7	-40.0	-7.7
qmj	Quality	8.6	6.2	8.6	-3.5	21.0
qmj_safety	Quality	-1.7	5.2	-1.7	-12.0	8.6
sale_bev	Quality	11.0	6.8	11.0	-2.2	24.3
seas_6_10an	Seasonality	8.6	2.6	8.6	<b>3.5</b>	<b>13.8</b>
seas_11_15an	Seasonality	7.2	3.1	7.1	<b>1.1</b>	<b>13.4</b>
rd_me	Size	8.7	3.3	8.7	<b>2.3</b>	<b>15.3</b>
rmax5_rvol_21d	Short-term reversal	6.7	2.4	6.7	<b>2.0</b>	<b>11.6</b>
fcf_me	Value	4.8	4.0	4.8	-2.9	12.7
div12m_me	Value	-1.3	5.4	-1.3	-11.8	9.3

## Summary

- Determining whether a return signal is a *true anomaly* requires knowledge of the true SDF.
- As a result, proving *non-spanning* is, in principle, impossible.
- In contrast, proving spanning only requires exhibiting an admissible factor set that prices the signal.
- Our analysis shows that, rather than an anomaly zoo, we are left with a small, well-behaved petting farm.

## Summary: way to do it

- Burden shifting:

$$H_{0,i} : a_i \notin \mathbf{w} \text{ (not spanned),} \quad H_{1,i} : a_i \in \mathbf{w} \text{ (spanned).}$$

- Model uncertainty:

- Integrate over all spanned/unspanned partitions of factors.
- Evidence is summarized by posterior model probabilities, not by a single benchmark test.

- Multiplicity:

- EFDR-controlled selection of spanned sets.

## Summary

- **FF6 stepwise scan:**
  - 126 anomalies classified as spanned.
  - 27 anomalies remain unspanned.
- **Investment relevance:**
  - FF6 augmented with 5–6 late-stage survivors delivers higher in-sample and out-of-sample tangency Sharpe ratios.
  - Cross-sectional pricing fit dominates that of size-matched fake anomaly sets.

## Summary

- Cross-benchmark robustness:
  - Union of spanned sets across FF6 and NewB6: 132 anomalies.
  - Residual unspanned set: 21 anomalies.
- SDF implications:
  - Eight factors (four from FF6, four from the survivor set) have strictly positive MPR credible intervals.